

## **RESEARCH ARTICLE**

# THE ROLE OF ENERGY CONSUMPTION IN CARBON DIOXIDE (CO<sub>2</sub>) EMISSION AND ECONOMIC GROWTH RELATIONSHIP IN NIGERIA

## Deekor<sup>1</sup>, LeeLee N.<sup>1</sup> and Aziaka Duabari S<sup>2</sup>

- 1. Department of Economics, Ignatius Ajuru University of Education, Port Harcourt, Nigeria.
- 2. Department of Mechanical Engineering, Ken-SaroWiwa Polytechnic, Bori, Rivers State Nigeria.

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*Key words:-*CO<sub>2</sub> Emission, Economic Growth, Energy Consumption, Nigeria, ARDL This paper uses the Autoregressive Distributed Lag(ARDL) method which allows for the combination of variables with mixed order of integration in a single regression model to the hypothesis that energy consumption matters in the carbon dioxide effect of economic growth. Using Nigerian dataset, our findings suggest that there is probable of a cointegrating relationship among the variables of interest. However, the estimation results refute the Environmental Kuznets Curve (EKC) in the context of the Nigerian economy. Essentially, we find the relationship between carbon dioxide (CO<sub>2</sub>) emissions and Gross Domestic Product (GDP) per capita to be monotonically increasing. We also find the response of CO<sub>2</sub> emission to energy consumption to vary for different energy-mix.

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## **Introduction:-**

Energy consumption in the process of production is considered as a precondition for the attainment of economic growth and sustainable development. It is indispensable for economic activity because all production and consumption activities are directly related to energy consumption. Future economic growth crucially depends on the long-term availability of energy from sources that are affordable, accessible, and environmentally friendly. However, the simultaneous rise in the production activities that generate economic growth cannot be in isolation of energy consumption, thus, motivating our hypothesis that the consumption of energy matters in the carbon dioxide( $CO_2$ ) and economic growth relationship. By using the transitive property in mathematics, we can infer that economic growth thrives on energy consumption which in turn drives carbon dioxide emissions (Bosupeng, 2016). Thus, energy consumption plays the dual role of providing the foundation for economic activity and human well-being as well as acting as the driving force for environmental degradation. Thus, the energy consumption –growth driven hypothesis has the potential to cause high carbon dioxide emitters, particularly for energy intensity economy such as Nigeria.

Partially due to their continuous expanding towards industrialization and urbanization stage of development, energy consumption and environmental degradation have continued to gain prominence in developing and energydependent economies such as Nigeria. Given that it is a growing economy, Nigeria has huge energy demands and energy requirements. As an oil-dependent economy, it is well established that all hydrocarbon extraction activities generate  $CO_2$  emissions. One particular by-product of crude oil production is associated with gas, the flaring of which generates large amounts of greenhouse gases (Total.com, 2018). On the net calorific value, Nigeria's economy is fueled by unclean and traditional energy, comprising 80.9 per cent of the total consumption. Cleaner and modern energy like gas and electricity comprised only a paltry amount of 11.1 per cent (Rapu et al., 2015). The sustainability of the energy systems in Nigeria is likely to be vulnerable if the anticipated energy crisis – in

## **Corresponding Author:- Deekor**

Address:- Department of Economics, Ignatius Ajuru University of Education, Port Harcourt, Nigeria.

particular, the electricity crisis and  $CO_2$  emissions issues – are not addressed appropriately. This is because the country is still highly dependent on fossil fuels such as oil and gas in its productive activities which also represents other main causes of carbon ( $CO_2$ ) emissions.

As part of the global initiative to reduce the emission of  $CO_2$  around the world, Nigeria has also adopted several environmental policies and strategies to reduceher share of the global emission of  $CO_2$ . However, understanding the cost and economic implications associated with such initiatives cannot be in isolation of evidence-based facts on the extent to which rise in the production activities that generate economic growth is also expected to increase  $CO_2$  emissions. Economic growth, energy consumption and carbon emissions are interrelated and, therefore, their relationship must be examined using an integrated framework to avoid misspecificationUsing the case of Nigerian economy, the focal point of this paper is to understand the extent to which productive activities relying on energy consumption matter for the country's share of carbon dioxide emissions. In addition to this introductory section, the rest of this paper is structured as follows: Section 2 reviews the empirical findings of the previous studies. Section 3 is the methodology, while section 4 discusses the data and offers some preliminary analyses. Section 5 presents the empirical results and discusses the findings while section 6 concludes the paper.

## Literature review:

The earlier literature analyzed the relationship between economic growth and environmental degradation according to the EKC analytic scheme and proposed an inverted U-shaped relationship between economic growth and environmental degradation (Grossman and Krueger, 1991; Stern et al., 1996; Ekins, 1997; Gani, 2012). Subsequently, scholars started to review empirical EKC studies (Dinda, 2004; Stern, 2004). In their pioneering work, Grossman and Krueger (1991) proposed an inverted U-shaped relationship between economic growth and environmental degradation. In recent time, however, empirical investigation on economic growth and environmental degradation has been extended to include the extent to which consumption of energy in the production activities accelerate the emission of  $CO_2$ . Notwithstanding, the intensity of research on the relationship between energy consumption, economic growth, and environmental pollution, the empirical evidence remains controversial and ambiguous to date(see, for example, Stern, 2004; Galeotti 2006; Ozturk et al., 2010;Belke et al., 2011; Halkos and Tzeremes 2011; Yang et al., 2012; Ahmad et al., 2012; Pirlogea and Cicea, 2012; Yali 2014;Esen, 2017; Taylor et al., 2018).

We also acknowledge the growth of the strand of empiricalliterature that captures energy consumption as a potential determinant of CO<sub>2</sub>.Lean and Smith (2009) for example, examined the causal relationship between carbon dioxide emissions and energy consumption through a panel vector error correction model for five ASEAN countries over the period 1980–2006. The long-run estimates indicate that there is a statistically significant positive association between energy consumption and emissions. In Iran, a one-way causal relationship from energy consumption (petroleum products and natural gas consumption) to CO<sub>2</sub> emission was found in the case of (Lotfalipour et al., 2010). In South Africa, Menyahand Rufael (2010) found a positive effect of CO<sub>2</sub> emissions on energy consumption. Similarly, Niu et al. (2011) show a positive relationship between energy consumption and  $_{CO2}$  emissions in eight Asian economies. Some studies considered otherforms of energy indicators. The studies by Tiwariet al. (2013); Shahbaz etal.(2013a), include coal consumption, the studyby Lotfalipour etal.(2010), includes fossil fuel consumption, the studyby Iwataetal.(2010), analyses the roleofnuclearenergy inFrance.Cowanetal.(2014), Farhani and Shahbaz (2014), Al-Mulali andOzturk(2015), Bento and Moutinho (2016), among others have also considered electricity power consumption as a proxy for energy consumption.

For Adewuyi and Awodumi (2017), it was argued that studies examining the relationship between energy consumption and economic growth without considering carbon emissions do not contribute much to the literature. Summarizing the inconsistency that has characterized findings of the existing studies on  $CO_2$  emissionsgrowth relationships, He (2009), argues that there is no one-fit-for-all growth-pollution relationship even when using the same estimation method. For example, Tamazian and Rao (2010), Taguchi (2012) and Gholami and Shafiee (2013) used GMM and they found inverted U, linear and N-shapes, respectively, although they used different samples of countries as case studies.

## The Model:

The EKC hypothesis postulates an inverted U-shaped relationship between  $CO_2$  emissions and per capita income: emissions per person increase up to a certain threshold level as per capita income goes up, after which they start to decrease (Dinda 2004; Müller-Fürstenberger and Wagner 2007; Kaika and Zervas, 2013). Following Storm and Mir

(1)

(2016), this study estimates a general reduced-form model in which  $CO_2$  emissions per capita is a polynomial cubic function (of degree three) of per capita income:

$$CO_{2t} = \beta_0 + \beta_1 Y_t + \beta_2 Y_t^2 + \beta_3 Y_t^3 + \beta_4 EC_t + \varepsilon_t$$

Where  $CO_{2t}$  denotes per capita emissions of carbon dioxide while  $Y_t$  denotes per capita income in real terms (measured at levels, quadratic and cubic form). The rapid growth of a country's economy may be caused by overdevelopment and inappropriate use of natural resources which will result in pollution and destruction of the environment and environmental crisis. The development and utilization of natural resources will increase the emissions of industrial pollutants, environmental degradation and resource depletion. Another main driver of carbon dioxide emissions is energy consumption given that the bulk of developing countries depend largely on fossil fuel consumption.

	Values of coefficients $\beta_i$	Relationship between income per capita (Y) and CO <sub>2</sub> emissions per capita
1	$\beta_1 = \beta_2 = \beta_3 = 0$	No relationship
2	$\beta_1 > 0$ and $\beta_2 = \beta_3 = 0$	A monotonically increasing or linear relationship
3	$\beta_1 < 0$ and $\beta_2 = \beta_3 = 0$	A monotonically decreasing relationship
4	$\beta_1 > 0, \beta_2 < 0, \beta_3 = 0$	An inverted-U-shaped relationship (KC)
5	$\beta_1 < 0, \beta_2 > 0, \beta_3 = 0$	A U-shaped relationship
6	$\beta_1 > \beta_2 < \beta_3 > 0$	An N-shaped relationship
7	$\beta_1 = \beta_2 = \beta_3 = 0$	An inverted-N-shaped relationship
8	$\beta_4 > 0$	The higher the choice of fossil fuels in energy consumption, the higher the carbon emission.
9	$\beta_5 < 0$	The higher the carbon dioxide emission, the lower the life expectancy.
10	$\beta_6 > 0$	Higher FDI inflows will require higher energy consumption which generates higher $CO_2$ emission
11	$\beta_7 > 0 \text{ or } \beta_7 < 0$	There is ambiguity in the effect of financial development and $CO_2$ emission.

Table 1:- Apriori expectation.

It can be seen that the EKC is only one of various possible numerical outcomes for equation (2), namely outcome 4 in Table 1, which occurs when we find that  $\beta_1 > 0$ ,  $\beta_2 < 0$  and  $\beta_3 = 0$ .

#### **Estimation technique:**

To examine the long-run relationship among the variables, we employ the use of the ARDL Bounds co-integration testing approach developed by Pesaran and Shin (1999) and further extended by Pesaran et al. (2001). The technique is largely preferred by economists and econometricians due to its flexibility to simultaneously accommodate variables with mixed order of integration in the same regression. Again, the Bounds test co-integration approach provides robust long-run estimates even in the presence of some endogenous variables in the model (Narayan, 2005). Finally, unlike the conventional techniques such as the Johannson co-integration approach, the bounds test is capable of giving robust results even when the sample size is small. Hence, these advantages make the adoption of the ARDL approach suitable in investigating the long-run impact of  $CO_2$  emission, energy consumption and economic growth in Nigeria. Using the bounds test approach, the following unrestricted error correction model will be estimated through the OLS method.

$$\Delta \ln CO_{2t} = \alpha + \sum_{i=1}^{p} \lambda_{1i} \Delta \ln CO_{2t-i} + \sum_{i=0}^{q} \lambda_{2i} \Delta \ln Y_{t-i} + \sum_{i=0}^{r} \lambda_{3i} \Delta \ln Y_{t-i}^{2} + \sum_{i=0}^{r} \lambda_{4i} \Delta \ln Y_{t-i}^{3} + \sum_{i=0}^{\nu} \lambda_{5i} \Delta \ln EC_{t-i} + \varphi_{1} \ln CO_{2t-1} + \varphi_{2} \ln Y_{t-1} + \varphi_{3} \ln Y_{t-i}^{2} + \varphi_{4} \ln Y_{t-i}^{3} + \varphi_{5} \ln EC_{t-1} + \varepsilon_{t}$$
(2)

where the parameters,  $\lambda_{mi}$  for m= 1, 2 . . . 5, represent the short-run dynamics in the model while the long-run relationships are given by  $\phi_i$ . To determine the long-run relationship between the regressand and regressors, the ARDL bounds test approach requires estimating equation (2) and restricting the parameters of the lag level (long run) variables to zero. Hence we test the null hypothesis (no co-integration)  $H_0$ :  $\phi_1 = \phi_2 = \phi_3 = \phi_4 = \phi_5 = 0$  against

the alternative hypothesis of co-integration. The hypothesis is tested using the F-test. The computed F-statistic is then compared with the Pesaran et al. (2001) asymptotic critical value bounds to ascertain the existence of a longrun relationship (co-integration). The null hypothesis of no co-integration is accepted if the computed F-statistic is less than the lower bounds and vice versa. The decision, however, remains inconclusive, if F-statistics lies between lower and upper critical bounds. Thus, in the event of a level relationship among the variables, the resulting long-run model can be estimated as:

$$\ln CO_{2t} = \phi_0 + \phi_1 \ln Y_{t-1} + \phi_2 \ln Y_{t-1}^2 + \phi_3 \ln Y_{t-1}^3 + \phi_4 \ln EC_{t-1} + v_t$$
(3)

The concluding step of the bounds test is to estimate the short-run elasticities which are obtained via the error correction framework represented by equation (4):

$$\Delta \ln CO_{2t} = \delta_1 + \xi_{ect} ECT_{t-1} + \sum_{i=1}^p \lambda_{1i} \Delta \ln CO_{2t-i} + \sum_{i=0}^q \lambda_{2i} \Delta \ln Y_{t-i} + \sum_{i=0}^r \lambda_{3i} \Delta \ln EC_{t-i} + \eta_t \quad (4)$$

where  $ECT_{t-1}$  is the error correction term while  $\xi_{ect}$  is the coefficient which captures the speed of adjustment of the model to its long-run equilibrium. In other words,  $\xi_{ect}$  captures the rate of correction at time t of deviation from the long-run equilibrium at time t-1.

## **Data and Preliminary Analysis:**

Based on the empirical specification, annual time series data will be collected for Nigeria between 1970 and 2016. Gross domestic product per capita at a constant price of 2005 expressed in US dollar serves as a proxy for real income per capita ( $Y_t$ ). Energy consumption (EC) is measured as fossil fuel energy consumption (ENC) as a percentage of the total. Nevertheless, the robustness of the energy consumption indicator will also be tested with energy use (ENU) measured as kg of oil equivalent per capita and electricity power consumption (EPC) as kWh per capita. The CO<sub>2</sub> emission is measured as CO<sub>2</sub> emissions (metric tons per capita). All the data are sourced from the World Bank World Development Indicator.

Reported in Table 2 are the statistical properties of the series including summary statistics and unit root testing results. Starting with the summary statistics in the (a) part of table 2, the mean statistic shows that average GDP per capita in Nigeria is 1.72\$ million per capita and 0.64 metric tons per capita for the  $CO_2$  emission series (GHC). All the series are negatively skewed but GDP per capita (GDPPC), while the kurtosis statistic is mostly platykurtic for all the series but ENC. Also, the null hypothesis of non-normality of the distribution of the series appears to be rejected for all the series but ENC. Concerning the unit root test, we consider both the conventional Augmented Dickey-Fuller (ADF) unit root test and modified version namely, Dickey-Fuller- Generalized Least-Squares (DF-GLS). The evidence mixed order of cointegration as reported in the (b) part of the table for both the ADF and DF-GLSfurther support our choice of estimation technique which allows for the combination of series with mixed order of cointegration.

Statistics	GDPPC	GHC	ENC	ENU	EPC				
Mean	1721.29	0.6359	17.7238	693.7958	89.9362				
Std. Dev.	431.31	0.1894	4.4124	55.5471	35.6967				
Skewness	0.4421	-0.0338	-1.5748	-0.4254	0.1134				
Kurtosis	1.9397	1.9974	4.6064	2.6318	2.2287				
J-Berra	3.7332	1.9777	24.4801	1.6829	1.2659				
	(0.1547)	(0.3720)	(0.0000)	(0.4311)	(0.5310)				

#### Table 2(a):- Summary Statistics.

#### Table 2(b): Unit Root Test

Variable	ADF		ADF-GLS					
	Level	First Difference	I(d)	Level	First Difference	I(d)		
GDPPC	-0.6504 <sup>b</sup>	-6.0590 <sup>b</sup> ***	I(1)	-0.7600 <sup>b</sup>	-5.1178 <sup>b</sup> ***	I(1)		
GHC	-2.2066 <sup>a</sup>	-6.8597a***	I(1)	-1.7192 <sup>b</sup> *	-	I(0)		
ENC	-2.4770 <sup>b</sup>	-5.7695 <sup>b</sup> ***	I(1)	-1.3730 <sup>b</sup>	-5.8937 <sup>b</sup> ***	I(1)		
ENU	-2.4633 <sup>b</sup>	-5.2402 <sup>b</sup> ***	I(1)	-1.6024 <sup>b</sup>	-5.6428 <sup>b</sup> ***	I(1)		
EPC	-3.3328 <sup>b</sup> *	-	I(0)	-2.3833 <sup>b</sup>	-9.2510 <sup>b</sup> ***	I(1)		

**Note:** The exogenous lags are selected based on Schwarz info criteria while \*\*\*\*, \*\*, \* imply that the series is stationary at 1%, 5% and 10% respectively. The null hypothesis is that an observable time series is not stationary (i.e. has unit root).



Figure 1:- Trends in CO<sub>2</sub> emission (GHC) and Economic Growth.

Figure 1 depicts possible co-movement between GHC and economic growth measured as  $CO_2$  emission and GDP per capita, respectively. Quite an interesting in the figure is the potential of positive and negative co-movements between the economic growth and  $CO_2$  emission. A cursory look at the figure, for example, shows that both the  $CO_2$  emission and economic growth appear to be moving in the same direction in the period between 1970 and 1987. The movement is, however, in the opposite direction for the period between 1986 and 2016. This though gives little or no statistical credence, yet it provides us with pre-information on the potential of the EKC hypothesis in the  $CO_2$  and economic growth relationship in Nigeria.

#### **Empirical Results:-**

Presented in Table 3 are empirical estimates from log-linear, log-quadratic, and log-cubic version of  $CO_2$  emission – economic growth functions. Starting with the Bound cointegration testing results, the decision on whether to reject the null hypothesis of no long-run relationship appears to be statistically indistinct in both squared and cubic with the F-statistic hovering between the upper and lower bounds of the critical values. The hypothesis of no cointegration is, however, significantly rejected when the GDP per capita is linearly expressed. Also, the fact that only the coefficient on GDP per capita appears to be statistically significant both in the linear and squared models thus suggesting that the variance in production related to  $CO_2$  emission is mainly linear in the case of the Nigerian economy. To put it differently, the estimated coefficient on the income variable (i.e. GDP per capita) has a positive sign, implying a linear relationship (monotonically increasing) among income and emissions. This by implication suggests that the EKC hypothesis does not hold for Nigeria over the period under consideration.

To examine the extent to which energy consumption matters in the  $CO_2$  –economic growth relationship, we further extend the preferred carbon emission function which in this case is the linear function to include the energy consumption variables that is under consideration. Contrary to our earlier finding, the Bound cointegration testing resultsseem to be suggesting that the null hypothesis of no long-run relationship holds for Model\_1 and Model\_2 where the energy consumption is measured as ENC and ENU but rejected in Model\_3 when energy consumption is measured as EPC. This by implication suggests that the probability of a long relationship between  $CO_2$  emission and economic growth in a model that controls for energy consumption might be sensitive to the measure of energy consumption that is under consideration. Consequently, we find little or no significant evidence that the  $CO_2$  emission is due to economic activity in the extended model, except when energy consumption in the model is measured as kg of oil equivalent per capita. This is quite expectant of the crude oil-producing economy. However, the negative sign on the coefficient on electricity power consumption (EPC) is an indication that renewable energy if initiated vial electricity consumption has the potential for reducing  $CO_2$  emission in Nigeria.

	Linear Model			Quadratic N	Quadratic Model			Cubic Model				
Short-Run	Coef	ficient	SE	T-stat	. Coefficient	SE	T-stat	. Coefficie	nt SE	T-stat.	,	
Constant	-1.57	/09*	0.8625	-1.821	2 -8.3035	4.7678	-1.741	5 -22.4593	20.70	78 -1.0845	5	
Trend	-0.00	)52**	0.0021	-2.426	7 -0.0030	0.0026	5 -1.149	7 -0.0030	0.002	6 -1.1475	5	
$\Delta \log(co_{2t-1})$	-0.29	937***	0.0897	-3.273	0 -0.3258***	0.0914	-3.564	0 -0.3257**	** 0.091	9 -3.5409	9	
$\Delta \log(y_t)$	0.2095* 0.1164 1.7		1.7997	1.1748*	0.6823	3 1.7217	3.3133	3.119	7 1.0620	)		
$\Delta \log(y_t^2)$					-1.58E-07	1.10E- 07	-1.435	-1.4351 -1.16E-06		0.8112	2	
$\Delta \log(y_t^3)$								2.37E-10	3.38E 10	- 0.7027		
ECM <sub>t</sub>	-0.29	937***	0.0818	-3.972	8 -3.5882***	0.0820	) -3.972	8 -0.3257**	** 0.080	1 -4.066	1	
Long-Run												
$\log(y_t)$	0.71	32**	0.3632	1.963	2 3.6060*	2.0248	1.780 8	10.1729	9.8149	1.0364		
$\log(y_t^2)$					-4.86E-07	3.33E- 07	- 1.460	-3.57E-06	4.49E- 06	-0.7937		
1 ( 3)							+	7 28E-10	1.06E-	0.6890		
$\log(y_t^*)$								7.201 10	09	0.0090		
Bound cointe	gratio	n testing	g result						_			
Level	0	f Line	ar Model		Quadratic N	Quadratic Model C			el	1		
Significance		F- stat	1(0)	<b>I</b> (1)	F-stat	1(0)	I(1)	F-stat	1(0)	I(1)		
10%			2.63	3.35		2.37	3.20		2.20	3.09		
5%		3.00	3.10	3.87	2.8760	2.79	3.67	2.4494	2.56	3.49		
1%		43	4.13	5.00		3.65	4.66		3.29	4.37		
Post-estimation	on res	ult										
			-		Linearity test	Aut	ocorrelati	on test	n test			
Model		Adj-R2	F-st	at.	Ramsey RESE	Г Q-S	tat.	Q <sup>2</sup> -Stat.	ARC	ARCH-LM test		
Linear		0.7203	39.6 *	423**	0.7561 (0.3896)	1.42	37 91)	1.4237 (0.491)	0.017	1 (0.9830)		
Quadratic		0.7272	30.9	968**	Not applicable	0.63	40	0.6340	0.066	0.0665 (0.9357)		
			*			(0.72	28)	(0.728)				
Cubic		0.7238	24.5	900**		0.29	92	0.2992	0.133	3 (0.8751)		
						(0.8)	01)	(0.801)				

Table 3:- ARDL estimates of carbon dioxide –GDP per capita nexus.

**Note:** The values in parenthesis represent the probability values for the various post estimation tests performance, while \*\*\*, \*\* and \* denote 1%, 5% and 10% level of significance.

Table 4:- ARDL Estimates for the role of energy consumption in CO<sub>2</sub> -GDP per capita Nexus.

	Model_1			Model_2			Model_3		
Short-Run	Coefficient	SE	Т-	Coefficient	SE	Т-	Coefficient	SE	Т-
			stat.			stat.			stat.

Constant	-1.4974	1.3128	-	-5.5226	6.5401	-	0.58	32	1.2180	0.4788		
T	0.0050	0.0024	1.1400	0.0007	0.00(1	0.8444	0.00	<u> </u>	0.0052	1.0114		
Irend	-0.0050	0.0034	-	-0.0087	0.0061	-	0.00	64	0.0053	1.2114		
		0.4050	1.4857			1.4336			0.0004			
$\Delta \log(co_{2t-1})$	-0.2895**	0.1073	-	-0.3254**	* 0.1043	-	-0.1	755*	0.0984	-		
21 17			2.6978			3.1203				1.7822		
$\Delta \log(y_t)$	0.2011	0.1623	1.2390	0.2420*	0.1288	1.8783	0.10	17	0.1192	0.8531		
$\Lambda \log(enc)$	-0.0007	0.0102	-									
			0.0750									
$\Delta \log(oil_t)$				0.5771	0.9466	0.6096						
$\Lambda \log(alaa)$							-0.34	547**	0 1486	-		
$\Delta \log(e l e c_t)$							0.5.	517	0.1100	2 3869		
ECM	_0 2895***	0.0806	_	_0 325/**	* 0.0889	_	-0.1	755***	0.0383	2.3007		
$ECM_t$	-0.2075	0.0000	3 5804	-0.3234	0.0007	3 6606	-0.1	155	0.0505	1 5736		
Long Dun			5.5694			5.0000				4.3730		
Long-Kun	0.0049	0 4 4 7 7	1 5510	0 7420**	0 22 47	2 2217	0.57	07	0 5796	1 0010		
$\log(y_t)$	0.6948	0.4477	1.5518	0.7438**	0.3347	2.2217	0.57	97	0.5786	1.0018		
log(enc.)	-0.0026	0.0360	-									
00000			0.0739									
$\log(oil_t)$				0.5771	0.9466	0.6096						
log(alac)							-2.02	208	1.7228	-		
$\log(e_t e_t)$										1.1729		
Bound cointegr	ration testing	result								111/2/		
Level of	Model_1			Model_2			Mod	lel_3				
Significance	F-stat	I(0)	<b>I</b> (1)	F-stat	I(0)	I(1)	F-sta	at	<b>I(0)</b>	I(1)		
10%		2.37	3.20		2.37	3.20			2.37	3.20		
5%	2.34	2.79	3 67	2.44	2.79	3 67	3.81		2.79	3 67		
1%		3.65	1 66		3.65	1.66	0.01		3.65	1 66		
Post ostimation	rocult	5.05	4.00		5.05	4.00			5.05	4.00		
1 Ost estimation	result	Madal	1		Model 2			Madal	2			
Adj $\mathbf{P}^2$		Model_	1		0.7161							
Auj-K E stat	lj-R <sup>2</sup> 0.7135		0.7161			0.7485						
<b>F-stat.</b> 29.0292***		0./155	***		20 2709**	*		21 101	<b>5</b> ***	34.4815***		
T-stat.		0.7135 29.0292	***		29.3798***	*		34.481	5***			
Ramsey RESE	ГTest	0.7135	***		29.3798*** 0.5621 (0.4	* 1578)		34.481 3.0288	5*** (0.0895)			
Ramsey RESE	Г Test	0.7135 29.0292 0.7582 ( 1.3079 (	*** (0.3891) (0.520)		29.3798*** 0.5621 (0.4 2.4812 (0.2	* 1578) 289)		34.4813 3.0288 0.3543	5*** (0.0895) (0.838)			
Ramsey RESE Q-stat. Q <sup>2</sup> -stat.	Г Test	0.7133 29.0292 0.7582 ( 1.3079 ( 1.3079 (	*** 0.3891) 0.520) 0.520)		29.3798*** 0.5621 (0.4 2.4812 (0.2 2.4812 (0.2	* 4578) 289) 289)		34.481 3.0288 0.3543 0.3543	5*** (0.0895) (0.838) (0.838)			
Ramsey RESE' Q-stat. Q <sup>2</sup> -stat. ARCH-LM	T Test	0.7135 29.0292 0.7582 ( 1.3079 ( 0.0116 (	**** (0.3891) (0.520) (0.520) (0.9884)		29.3798*** 0.5621 (0.4 2.4812 (0.2 2.4812 (0.2 0.0957 (0.9	* 1578) 289) 289) 2089)		34.4813 3.0288 0.3543 0.3543 0.0879	5*** (0.0895) (0.838) (0.838) (0.9060)			
Ramsey RESE Q-stat. Q <sup>2</sup> -stat. ARCH-LM	T Test	0.7135 29.0292 0.7582 ( 1.3079 ( 1.3079 ( 0.0116 (	**** (0.3891) (0.520) (0.520) (0.9884)		29.3798*** 0.5621 (0.4 2.4812 (0.2 2.4812 (0.2 0.0957 (0.5	* 1578) 289) 289) 0089)		34.481: 3.0288 0.3543 0.3543 0.0879	5*** (0.0895) (0.838) (0.838) (0.9060)			
Ramsey RESE Q-stat. Q <sup>2</sup> -stat. ARCH-LM	<b>T Test</b> URE	0.7135 29.0292 0.7582 ( 1.3079 ( 1.3079 ( 0.0116 (	*** 0.3891) 0.520) 0.520) 0.9884)		29.3798*** 0.5621 (0.4 2.4812 (0.2 2.4812 (0.2 0.0957 (0.5	* 1578) 289) 289) 0089)		34.481: 3.0288 0.3543 0.3543 0.0879	5*** (0.0895) (0.838) (0.838) (0.9060)			
Ramsey RESE' Q-stat. Q <sup>2</sup> -stat. ARCH-LM NOMENCLATI ARDL - Autore	<b>T Test</b> URE gressive Distri	0.7135 29.0292 0.7582 ( 1.3079 ( 0.0116 (	*** 0.3891) 0.520) 0.520) 0.9884) ,		29.3798*** 0.5621 (0.4 2.4812 (0.2 2.4812 (0.2 0.0957 (0.9	* 1578) 289) 289) 2089)		34.481: 3.0288 0.3543 0.3543 0.0879	5*** (0.0895) (0.838) (0.838) (0.9060)			
Ramsey RESE' Q-stat. Q <sup>2</sup> -stat. ARCH-LM NOMENCLAT ARDL - Autore EKC Environr	<b>T Test</b> URE gressive Distri nental Kuznets	0.7135 29.0292 0.7582 ( 1.3079 ( 0.0116 ( buted Lag	*** 0.3891) 0.520) 0.520) 0.9884) ,		29.3798*** 0.5621 (0.4 2.4812 (0.2 2.4812 (0.2 0.0957 (0.9	* 1578) 289) 289) 2089)		34.481: 3.0288 0.3543 0.3543 0.0879	5**** (0.0895) (0.838) (0.9060)			

ENC-- Energy consumption

ENU --Energy use

EPC --Electricity power consumption.

GLS- Generalized Least- Squares. GMM- Gaussian Mixture Model

GMM- Gaussian Mixture Model

**Note:** The energy consumption in model\_1 is represented by fossil fuel energy consumption measured as a percentage ofthe total, log of kg of oil equivalent per capita in model\_2, and log of electricity power consumption in kWh per capita in model\_3. Value in parenthesis represents the probability values for the various post estimation tests performance, while \*\*\*, \*\* and \* denote 1%, 5% and 10% level of significance.

## **Concluding Remark:-**

This paper uses Nigerian dataset to examine carbon dioxide emission effects of economic growth while accounting for the role of energy consumption. To capture both the short and long-run dynamics of the relationship simultaneously, we explore the ARDL method which also allows for the combination of the variable with mixed order of integration in the same regression. Our findings though suggest that there is probable of a cointegrating relationship among the variables of interest, the estimation results yet suggest the EKC hypothesis does not hold in the case of the Nigerian economy. Essentially, we find the relationship between  $CO_2$  emissions and GDP per capita to be monotonically increasing. We also find the response of  $CO_2$  emission to energy consumption to vary for different energy-mix. For instance, while energy consumption measured as kg of oil equivalent per capita is capable of causing increasing GHC, our finding of a negative sign on the coefficient on electricity power consumption (EPC) seems to be suggesting that renewable energy if initiated vial electricity consumption has the potential for reducing  $CO_2$  emission in Nigeria.

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